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Assimilation of Immigrants in Germany

Evidence from an Earnings Frontier Approach

by

Günter Lang

Beitrag Nr. 156

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Assimilation of Immigrants in Germany

Evidence from an Earnings Frontier Approach

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Abstract

This study uses the concept of stochastic frontiers for testing the assimilation hypothesis in the West-German labor market. Estimation of a potential rather than an average earnings function increases the explanatory power of the human capital approach and allows for the determination of individual levels of income efficiency. The results imply that immigrants from East Germany and Eastern Europe exhibit a significantly steeper earnings profile than natives, clearly supporting the assimilation hypothesis for this group. Ten years after arrival immigrants reach "wage parity" with West Germans. Surprisingly, only small differences could be observed with regard to the question of earnings efficiency. On an average, inhabitants as well as immigrants transformed about 75% to 80% of their potential income into actual earnings. The reasons for the individually diverging efficiency ratios are not well understood, with the market power of small firms against their employees being identified as the most important source.

JEL classification: J31, J61

Keywords: Immigration, Assimilation, Earnings Frontier

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Introduction

Worldwide immigrant flows have dramatically increased since the beginning of the eighties. Many of these immigrants have chosen a non-traditional destination, like Germany or France, for their new home country. As a result, more than 2 mio persons from Eastern Europe (mainly Russia, Poland and Romania) arrived in Germany. Because of their ethnic extraction and in contrast to guest workers or refugees, these people have legitimate claim for permanent residence in Germany. Most of the East European immigrants moved to the western part of Germany, where additional 1.5 mio East Germans arrived. Ignoring guest workers and refugees, which may be viewed as temporary phenomena, more than 5% of the West German population is foreign born and migrated during the last 15 years. Not very surprising, this immigration process is accompanied by a heated political discussion.

In this paper, the assimilation hypothesis proposed by Chiswick (1978) is tested for these East European immigrants. Following Chiswick, foreign borns have lower wages than inhabitants immediately after immigration, but catch up in subsequent periods. The political importance of this question is evident: If the assimilation hypothesis is true and immigrants show an outperforming labor market success, their high productivity should substantially contribute to enhanced growth in the host country. Durable low wages are more difficult to interpret, however: On the one side, low earnings could be the result of discrimination against inhabitants, which raises important political and economic questions. On the other hand, this outcome could indicate a permanent productivity gap relative to natives, implying a long-term burden for the public welfare system.

Existing empirical work on the assimilation feature has concentrated on the US (for an overview see *Borjas*, 1994, Chiswick, 1994). For Germany, there exist some work for guest workers (e.g. Licht and Steiner, 1993, Dustmannn, 1993, 1996, Pischke, 1993), whereas the large immigrant flows from Eastern Europe are hardly analyzed (for some newer contributions see *Bauer and Zimmermann*, 1995, Schulz and Seiring, 1994). The reason for this deficit are insufficiencies in the database, which could be overcome with the latest supplement of the Socio-Economic Panel (see Schupp and Wagner, 1995, for a description). For 1994, this supplement consists of about 250 households, which immigrated from Eastern Europe (incl. East Germany) and show a German extraction.

To enhance the quality of the earnings estimates, not a classical "average" function, but a frontier technique is used. The estimation of frontier income functions is based on models which were originally introduced for production technologies (*Farell*, 1957). In

that framework, firms are allowed to produce inefficient. This means that not the maximum possible output is produced with a given set of inputs, but a lower volume - with varying degrees of inefficiencies between firms (*Fried et al., 1993*, give a survey about estimation methods). Imperfect markets allow for the permanent existence of such deviations from the frontier function. Transferred to the labor market, individual-specific information deficiencies, immobility or discrimination could prevent workers from realizing maximum possible earnings. Analogous to a production frontier, the earning frontier gives the highest potential income associated with varying amounts of human capital inputs.

The estimation of production or cost frontiers has become increasingly popular during the last years, especially for the analysis of financial institutions (for a survey see Berger and Humphrey, 1997). Within the field of labor economics, this approach has been used for some US studies (Herzog et al., 1985, Hofler and Polachek, 1985, Robinson and Wunnava, 1989, Hunt-McCool and Warren, 1993). For the immigration problem there exists a study of Daneshvary et al. (1992), who define assimilation as a decrease in the difference between actual and potential earnings and test this hypothesis for the US immigration market.

The following section contains the stochastic specification of the employed earning frontiers and the estimation methods. In section III the dataset is described. Furthermore, separate frontier earnings estimates for East European immigrants and inhabitants are presented. From these estimates, conclusions on the assimilation hypothesis can be drawn. Section IV analyzes deviations from the frontier in more detail, especially by comparing income efficiency from inhabitants and foreigners. The relationship between individual degrees of earning efficiency and some hypothetical sources like marital status or firm size is presented, too. Finally, section V sums up.

Specification

Basic assumption is the existence of a semilogarithmic human capital production function of the extended *Mincer* (1974)-type

$$\ln E_i = \alpha + \beta^i X_i + \varepsilon_i \qquad i = 1,...,n \tag{1}$$

where E_i denotes earnings of individual i, X is a vector of socio-economic characteristics, α and β are fixed but unknown parameters, and ε is the error term. Equation (1) assumes that wages are systematically dependent on the endowment with human capital X, which is proxied by the factors age, schooling, training, work experience and times of

unemployment. As better the endowment with human capital, as higher the productivity and therefore the market wage.

The income equation (1) is estimated separately for immigrants and inhabitants in Germany. To test the assimilation hypothesis proposed by *Chiswick (1978)*, earnings of immigrants are allowed to be additionally influenced by the term "years since migrations". The assimilation theory is supported if the empirical results show higher starting wages for natives, but a positive coefficient for the migration variable. This is due to the assumption that migration depresses the worth of human capital, which is compensated during the next years by factors like continuous training or increasing language skills, however. After a certain time and because of a high motivation, immigrants may even overtake earning levels of inhabitants. Therefore one has to compare the wage function of inhabitants with that of immigrants at different levels of "years since migration".

In this paper special interest is taken on the specification of the error term ε . The idea is to estimate a "frontier earnings function" for inhabitants as well as for foreigners, which gives the maximum earning one can receive with his given bundle of socio-economic characteristics. For this purpose, frontier estimation techniques, which can be divided into Data Envelopment Analysis (DEA) and econometric approaches, have to be used. Because of the parametric structure of the earnings function (1), the econometric frontier approach is appropriate for the estimation process. \exists conometric techniques can again be subdivided into deterministic and stochastic procedures, with the latter being used in this study because of its higher flexibility and the diminished outliers problem (for an overview see *Greene*, 1993).

The stochastic frontier approach can be characterized by an error term ε , which consists from two parts: $\varepsilon_i = v_i - u_i$. Standard random error v_i is represented by $v_i \sim N(0, \sigma_v^2)$, whereas u_i reflects labor market inefficiency of an specific person i. This second part u_i is restricted to be non-negative, because otherwise one would be allowed to earn more than his potential (maximum) earning. Characterized within the econometric framework, the deterministic part of the earnings function $\alpha + \beta' X_i$ defines an envelope associated with specific amounts of human capital. Since individuals can earn less because of information deficiencies, immobility and so on $(u_i \ge 0)$, the composite error term ε_i has a nonzero mean which reflects the systematic deviation of actual wage from the hypothetical norm.

To estimate the parameters of the underlying function, the stochastic distribution of the inefficiency term u_i has to be specified. Two of the most popular assumptions are the half normal distribution and the exponential distribution of u_i , introduced by Aigner et

$$EFF_{i} = \frac{\exp(\alpha + \beta' X_{i} - u_{i})}{\exp(\alpha + \beta' X_{i})} = \exp(-u_{i})$$
(6)

 EFF_i is restricted to the interval [0,1], with the upper boundary representing a worker who transforms his human capital endowment perfectly into market income. A value of less than one - e.g. 0.90 - indicates that the underlying employee actually earns 90% of this potential income.

The knowledge of individual values EFF_i raises the question for the sources behind these different levels in earnings efficiency. Therefore in a second step a simple regression was run, explaining EFF_i by some individual and market specific factors: ²

$$EFF_{i} = \delta_{0} + \delta_{1} * MARRIAGE + \delta_{2} * CHILD + \delta_{3} * SMALL + \delta_{4} * MEDIUM + \varepsilon_{i}^{1}$$
 (7)

MARRIAGE gives the marital status, CHILD the number of children living in the house-hold of individual i. To be married and educating children are important determinants of immobility, therefore the influence of these factors is expected to be negative. The dummy variables SMALL and MEDIUM take the value one if the firm, where worker i is employed, has less than 20 respectively between 20 and 2000 employees. Large firms with more than 2000 employees are the reference scenario. If unions are more powerful in big firms, δ_3 and δ_4 should have a negative sign. ε_i^1 represents a regular error term. Again, the estimation of (7) was run separately for inhabitants and immigrants.

Earnings Estimates and Assimilation

This paper strongly profited from the latest supplement to the German "socio-economic panel" (GSOEP), where a random sample of immigrants was interviewed. The bulk of this immigration sample consists of East Germans and East Europeans with a German extraction, who migrated to Western Germany since 1984. In contrast to guest workers, these persons can be considered as permanent immigrants without a return motivation. There are three arguments for this assumption:

 Economic, social and political conditions in many East European countries are not attractive for the German minority. This is especially true for Russia, which is the most important source of East European immigrants.

Alternatively one could specify a model where the level of inefficiency and some potential determinants of this inefficiency are entering the estimation process jointly (Reifschneider and Stevenson, 1991).

Table 1:
Description of the dataset

		Mean Value		
Variable	Definition	Inhabitants	Immigrants	
EARNING	gross DM-income from wages/salaries divided by hours of work	28.92	20.88	
AGE	age of person in years	41.50	39.43	
SCHOOL	years of schooling	10.67	9.91	
TRAINING	equals 1 if employee received occupa- tional training or university degree; 0 otherwise	0.94	0.85	
EXPERIENCE	actual full-time work experience in years (without time of apprenticeships)	19.99	16.67	
JOBLESS	time of unemployment in years	0.55	0.59	
IMMIGRATION	years since immigration to West-Germany	-	4.85	
MARRIED	equals 1 if married; 0 otherwise	0.82	0.93	
CHILD	number of children living in household	0.80	0.45	
SMALL	equals 1 if firm has less than 20 employees, 0 otherwise	0.16	0.27	
MEDIUM	equals 1 if firm has more than 20 and less than 2000 employees, 0 otherwise	0.47	0.58	
	Number of observations	1099	182	

Male employees only; all data for 1994. Source: German SOEP; own calculations.

- For East Germans, there may exist a return motive if economic conditions catch up
 with West Germany. The assimilation process should not be influenced, however,
 because country specific labor market conditions in the unified Germany will have
 disappeared at this time.
- From a formal point of view, the German extraction ensures the right for a German passport and permanent stay. No further residence permit is needed.

To test the assimilation hypothesis, the immigrant sample was compared to a sample of West-German inhabitants. To reduce distortions from self-selection, extreme outliers or measurement errors, the research population was confined to a subsample of males between 18 and 65 years of age. Furthermore, all self-employed or persons with missing information were excluded. In *Table 1* the variables and the datasets are described,

clearly showing a considerable higher income of inhabitants in comparison to immigrants.

To analyze this earning structure in more detail, income functions are separately estimated for both research populations (likelihood functions (2) and (3)). The wage function (1) was specified as

$$\ln E_{i} = \alpha + \beta_{1} * AGE + \beta_{2} * SCHOOL + \beta_{3} * TRAINING + \beta_{4} * EXPERIENCE + \beta_{5} * EXPERIENCE^{2} + \beta_{6} * JOBLESS + \beta_{7} * IMMIGRATION + \varepsilon_{i}$$
(8)

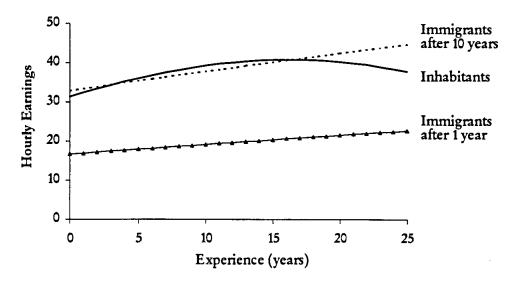
with $\beta_7 = 0$ for inhabitants.³ This functional form allows for concavity in the experience variable. No regressors like firm size, marital status and so on were added, because these factors aren't considered as components of the human capital stock. However, as equation (7) shows, these variables are allowed to influence the wedge between potential and actual earnings.

The results of the Maximum-Likelihood estimation and - for comparison - of the standard OLS-method are presented in *Table A- 2* in the Appendix. The iteration process turned out to be stable and converging. Most of the estimated parameters are statistically significant; exceptions are the training-variable for natives and the squared experience term as well as the jobless parameter for immigrants. Likelihood-ratio tests clearly reject the hypothesis that only random error exists ($\lambda = 0$ for the half-normal assumption; $\theta = 0$ for the exponential case). The advantage of the frontier approach is evident since the stochastic frontier approach reduces the residual variance σ_{ν}^2 by about 50% (natives) respectively 60-70% (immigrants).

Although all parameter signs are independent from the estimation method, remarkable differences between parameter values appear. These discrepancies are relative small between the half-normal and the exponential distribution, whereas the differences between the frontier and OLS model are more important. One additional year of schooling, for example, increases immigrant wages by 1.7% in the OLS case, but by more than 2.8% with the half-normal frontier estimation. Similarly, return on training decreases from 25% (OLS) to 16% with an exponential distributed inefficiency term.

To control for differences between East Germans and Eastern Europeans, additional estimations were run including a Dummy for East Germans. Neither for the earnings function (8) nor for the efficiency equation (7) could be found any significant influence. Furthermore, the parameter values turned out to be very stable against this modification.

Figure 1: Frontier Earning Functions of Inhabitants and Foreigners



Simulation based on human capital stock of a typical inhabitant (41.5 years old, 10.7 years of schooling, 0.94 received training, 0.55 years of unemployment).

However, with the main focus of this paper, differences between inhabitants and immigrants are of greater interest. As the parameter estimates show, return on schooling is greater for inhabitants (6.7% versus 2.8%), whereas training has a stronger effect for migrants (less than 6% versus more than 16%). The AGE variable even turns the sign from significantly positive for West-Germans to significantly negative for East Europeans. In the human-capital context, the negative sign is more in line with expectations, because older people are more likely to be ill and have less physical strength. The wage-increasing effect of getting older - accumulating job experience - is explicitly considered by the experience variable. Finally, the income of natives depends more heavily on individual skills than this could be observed for immigrants, as the constant term is higher for the second group.

Most important, however, the results strongly support the assimilation hypothesis in the sense of Chiswick (1978). This implication can be illustrated by Figure 1, where frontier earnings of a "typical" native employee is plotted against frontier earnings of immigrants. Equipped with an identical human capital stock as the native colleague, the immigrant has a potential income of about 60% relative to the inhabitant and immediately after arrival. The earnings profile turns out to be much steeper in the subsequent years, however: Ten years after arrival, the differences between natives and former East Europeans have disappeared. Former East Europeans have perfectly assimilated within a relatively short time period, at least in terms of income. If immigrants would even

overtake nationals after this period is unclear, because the sample only covers individuals who crossed the border after 1984. An extrapolation beyond this time horizon would therefore be speculative.

Earning Efficiency

Given the parameter estimates in Table A- 2, the underlying distribution of earnings inefficiency can be determined. For example, the density function of u in the half-normal case⁴ can be expressed as

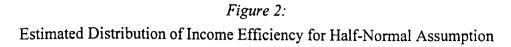
$$d(u) = \begin{cases} \frac{\sqrt{2}}{\sqrt{\pi}\sigma_u} \exp\left[-0.5\left(\frac{u}{\sigma_u}\right)^2\right] & \text{if } u \ge 0\\ 0 & \text{else} \end{cases}$$
 (9)

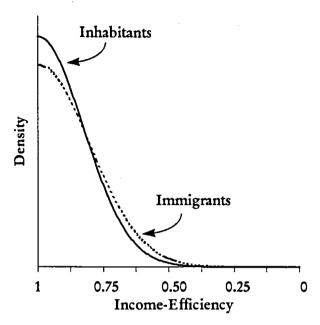
As can be seen from Figure 2, where the half-normal case is illustrated, the estimated differences between natives and foreigners are small⁵. Given the distribution in (9), the mean value of u is $\sigma_u \sqrt{2/\pi}$, which produces 0.280 for inhabitants and 0.319 for immigrants. Transformed to efficiency values EFF_i , the results are 75.6% versus 72.7%. This implies that immigrants are somewhat more inefficient in transforming human capital to market income, but the discrepancy turns out to be of minor interest.

From the earnings frontier estimates, individual values ε_i and u_i can also be determined for all natives and immigrants within the research population (equations (4) respectively (5)). Table A-1 in the appendix provides the distribution of the expected efficiency values, confirming the above mentioned results. A slightly more heterogeneous appearance of immigrants notwithstanding, the distances between estimated and potential income are very similar for both groups. Average XEFF values are close to the forecast from the distribution function. The ranking coefficient between the exponential and the half-normal model is at 0.99, supporting the reliability of the results. Actually, the major difference is the trivial more pessimistic assessment of the half-normal model.

For the exponential assumption, the distribution is $\theta \exp[-\theta u]$, $u \ge 0$.

⁵ A very similar picture arrives for the exponential case.





Income efficiency calculated as exp(-u).

In a second step, the relationship between wage inefficiency and its potential sources is estimated (equation (7)). Table 2 shows the regression results for both groups and the half-normal as well as the exponential distribution. As can be seen, to be employed at small firms with less than 20 employees significantly increases the wedge between potential and realized earning. In that case, the EFF_i value is estimated to decrease by about 8% (natives) respectively 9%-11% (immigrants). This general picture is in line with other studies⁶ and may be due to a smaller influence of unions, for example. Interestingly, the wedge between medium sized (20-2000 employees) and large firms (>2000 employees) is at just 2% and significant only for the native group.

Some open questions remain with regard to the marital status and the number of children: To be married should have a negative effect on mobility and therefore income efficiency, but this expectation doesn't find empirical support. Indeed, native married males can even approach their individual frontier by about 4%. The parameter for immigrants is insignificant, but positive, too. The existence of children in one's household

For example, Althammer and Wenzler (1996) and Bauer and Zimmermann (1995) also show a positive influence of firm size on earnings. Notice, however, that both studies use firm size as explaining variable within the earnings equation context, whereas in this paper firm size is explaining the difference between potential and realized earning. Only direct human capital measures are entering the earning equation (8).

also should decrease mobility and drive a wedge between potential and realized income, but this could be observed only for the immigrant population.

Table 2: Explanation of Earnings Efficiency

	Inhabitants		Immigrants	
	Half-Normal	Exponential	Half-Normal	Exponential
CONST	0.759***	0.803***	0.768***	0.838***
MARRIED	0.032***	0.049***	0.044	0.032
CHILD	0.005*	0.004	-0.051***	-0.046***
SMALL	-0.077***	-0.076***	-0.105***	-0.092***
MEDIUM	-0.019***	-0.016**	-0.022	-0.019
R²	0.089	0.112	0.175	0.168

Dependent variable is EFF_i , *, ** and *** represent a significance level of 90%, 95% and 99%, respectively (two-sided).

Finally, it should be noted that equation (7) can explain only a small part of the earnings efficiency differences. The determination coefficient takes values of about 10% for West-Germans respectively 15% for East Germans, which leaves a lot of room for further speculation about the reasons for diverging levels of income efficiency.

Conclusions

The 1980's were characterized by a resurgence of immigration to West-Germany, accompanied by a change in the origin mix from Southern Europe (Turkey, Yugoslavia, Italy, Spain, Greece) to Eastern Europe. These new immigrants are Germans from former East Germany or the descendants of German settlers in Romania and Russia. Not surprising, these huge immigration waves were accompanied by a political debate over the quantity and types of migrants, which can be adopted.

This paper is following recent research in population economics, assessing the labor market performance of these immigrants. On the basis of a stochastic earnings frontier, income profiles of East Europeans migrants are compared with those of inhabitants. The estimation of earnings frontiers allows for quantifying the maximal income which can be earned with a given level of human capital. Furthermore, individual differences between potential and actual income are estimated and put into relationship to some hypothetical factors of influence.

One of the most important results of this study is the empirical support for the assimilation process. East Europeans with German extraction, who count for a large part of the newly immigrated persons, show a significantly steeper earnings profile than natives. Starting with considerably lower wages immediately after arrival, only ten years later immigrants have adjusted to natives. Notice, however, that there are substantial differences in the parameters of the potential earning function: For example, return on schooling is greater for inhabitants, whereas apprenticeships or university degrees turned out to be more positive for migrants.

The observed evidence in favor of the assimilation hypothesis is in sharp contrast to the group of guest workers (Turks, Yugoslavs, Italians, ...). For these persons, who typically live in Germany for 25 or more years, no assimilation process could be observed (*Licht and Steiner*, 1993, *Pischke*, 1993). Of course, further research is necessary to confirm these differences, but obviously an immigration policy which is in favor of cultural neighbors should be preferred.

With regard to the wedge between potential and actual income, the stochastic frontier estimations are in line with other studies analyzing labor markets (see e.g. Daneshvary et al., 1992, Hunt-McCool, 1993). On an average about 75%-80% of the potential income can be realized as market earning. The difference between natives and East Europeans amounts to 2%, which seems insignificant. This last result is even more surprising than the rapid assimilation process, because newly arrived job searchers should have less information about the host countries labor market than natives.

As for the explaining factors of this wedge between potential and actual wage, to be employed at a small firm is clearly negative for workers. The market power of small firms seems sufficient to pay their employees beyond the value of their human capital endowment and therefore below their productivity. Somewhat surprising is the efficiency-increasing influence from the marital status, where married persons approach further to their earnings frontier than non-married. The opposite sign was expected because of a diminished mobility. Children living in one's household lowers income efficiency only for immigrants. In general, only a small part of earning inefficiency can be explained, leaving room for further speculation.

Summing up, the positive labor market performance of immigrants from Eastern Europe is providing support for further reception of these types of migrants. They could substantially contribute to higher growth rates and help stabilizing the welfare system. The cultural and ethnic similarity to native Germans are further reasons towards a corresponding immigration policy. Some doubts remain, however, if the success of the for-

mer immigrants can be extrapolated to the future: *Borjas (1985)* supposes for the United States that the quality of immigrants has deteriorated over time, and assimilation is much less important for earnings growth than cross-sectional comparisons indicate. He pointed out that cross-sectional studies may be biased because they don't disentangle intergenerational mobility from assimilation. With regard to Germany, a similar process could take place as the consequence of negative self-selection: Motivated by generous welfare transfers and high tax rates, not the best but the less successful or older persons are attracted to migrate to Germany. Policy should be aware of this problem and look for appropriate solutions.

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Appendix

Table A- 1
Distribution of Earnings Efficiency

	Inhabitants		Immigrants	
	Half-Normal	Exponential	Half-Normal	Exponential
< 0.50	3.0%	2.7%	6.6%	4.4%
0.50 - 0.60	4.3%	2.1%	8.8%	2.8%
0.60 - 0.70	43.1%	4.5%	14.3%	8.2%
0.70 - 0.80	36.5%	15.6%	26.9%	14.8%
0.80 - 0.90	37.9%	54.9%	33.5%	47.8%
0.90 - 1	5.3%	20.2%	9.9%	22.0%
verage Efficiency	0.768	0.826	0.744	0.812

Earnings Efficiency estimated as $EFF_i = \exp(-u_i)$

Table A- 2
Estimation Results

	Inhabitants			Immigrants		
	OLS	ML - Half Normal	ML - Exponential	OLS	ML - Half Normal	ML - Exponential
CONST	1.379 (16.524)	1.708 (20.740)	1.697 (21.564)	2.453 (12.08)	2.734 (15.275)	2.680 (15.353)
AGE	0.025 (8.815)	0.025 (8.475)	0.024 (8.169)	-0.010 (-1.845)	-0.011 (-2.159)	-0.011 (-2.181)
SCHOOL	0.062 (9.212)	0.065 (9.713)	0.065 (10.152)	0.017 (1.172)	0.028 (2.115)	0.025 (1.967)
TRAINING	0.050 (1.281)	0.048 (1.288)	0.055 (1.527)	0.225 (3.517)	0.174 (3.183)	0.155 (2.742)
EXPERIENCE	0.039 (9.309)	0.032 (7.489)	0.028 (6.583)	0.016 (2.071)	0.015 (2.181)	0.016 (2.403)
EXPERIENCE^2 *10²	-0.113 (-16.450)	-0.098 (-14.204)	-0.088 (-12.900)	-0.013 (-0.681)	-0.009 (-0.557)	-0.012 (-0.732)
JOBLESS	-0.041 (-6.452)	-0.040 (-6.635)	-0.039 (-6.635)	-0.006 (0.276)	-0.006 (-0.312)	-0.004 (-0.193)
IMMIGRATION	-	-	-	0.073 (4.592)	0.075 (5.635)	0.073 (5.464)
R²	0.449			0.221		
$\sigma_{\!\scriptscriptstyle u}^{\scriptscriptstyle 2}$	0.089	(9.463)	0.046 (12.868)	0.084	0.023 (2.839)	0.033 (4.214)
λ		2.799 (4.723)			6.936 (1.908)	
θ			4.940 (15.139)			4.484 (6.702)
observations	1099	1099	1099	182	182	182

Dependent variable is ln of hourly (gross) earnings. t-ratios in parentheses.

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