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Dietary fat and breast cancer risk in the European Prospective Investigation into Cancer and Nutrition^{1–3}

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ABSTRACT

Background: Epidemiologic studies have produced conflicting results with respect to an association of dietary fat with breast cancer.

Objective: We aimed to investigate the association between fat consumption and breast cancer.

Design: We prospectively investigated fat consumption in a large ($n = 319\,826$), geographically and culturally heterogeneous cohort of European women enrolled in the European Prospective Investigation into Cancer and Nutrition who completed a dietary questionnaire. After a mean of 8.8 y of follow-up, 7119 women developed breast cancer. Cox proportional hazard models, stratified by age and center and adjusted for energy intake and confounders, were used to estimate hazard ratios (HRs) for breast cancer.

Results: An association between high saturated fat intake and greater breast cancer risk was found [HR = 1.13 (95% CI: 1.00, 1.27; P for trend = 0.038) for the highest quintile of saturated fat intake compared with the lowest quintile: 1.02 (1.00, 1.04) for a 20% increase in saturated fat consumption (continuous variable)]. No significant association of breast cancer with total, monounsaturated, or polyunsaturated fat was found, although trends were for a direct association of risk with monounsaturated fat and an inverse association with polyunsaturated fat. In menopausal women, the positive association with saturated fat was confined to nonusers of hormone therapy at baseline [1.21 (0.99, 1.48) for the highest quintile compared with the lowest quintile; P for trend = 0.044; and 1.03 (1.00, 1.07) for a 20% increase in saturated fat as a continuous variable].

Conclusions: Evidence indicates a weak positive association between saturated fat intake and breast cancer risk. This association was more pronounced for postmenopausal women who never used hormone therapy. *Am J Clin Nutr* 2008;88:1304–12.

INTRODUCTION

Breast cancer is the most common cancer affecting women worldwide. For women living in low-risk countries, the risk of developing breast cancer increases upon immigration to a high-risk country (1), which suggests that this cancer is influenced by modifiable lifestyle or environmental factors.

Epidemiologic studies have produced conflicting results regarding an association of dietary fat with breast cancer.

Although strong associations between high fat intake and greater breast cancer incidence have been found in international correlation studies (2) and animal studies (3–6), most case-control studies indicate only a weak association (7), and prospective cohort studies have usually shown little (8) or no (9–12) association. A recent prospective study on postmenopausal US women found that dietary fat intake was weakly but significantly associated with the risk of invasive breast cancer, whereas intakes of saturated, monounsaturated, and polyunsaturated fat each were significantly related to breast cancer risk (13). A pooled analysis of data from cohort studies found a weak increase in breast cancer risk when carbohydrates were “substituted” by saturated fats in an isocaloric diet (14), and a meta-analysis of 45 epidemiologic studies (14 cohort studies and 31 case-control studies) found a weak but statistically significant relation between high fat intake and greater breast cancer risk (15). It is interesting that fat intake was significantly associated with breast cancer risk in European but not

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in Canadian or US studies, possibly in relation to greater variation in dietary fat intake among European than North American women (15). The recently published results of the randomized Women's Health Initiative found that low-fat diets were associated with a 9% reduction in the risk of invasive breast cancer (16).

Conflicting results with respect to monounsaturated fat have also been reported. Some cohort studies reported a protective effect of monounsaturated fat intake on breast cancer risk (12, 17–19), whereas others reported no effect or even a positive association with breast cancer (9, 11, 13–15, 20). The conflicting results of analytic epidemiologic studies are likely to be due to the known difficulties in obtaining precise estimates of intakes of various types of fat (21) and also to the limited heterogeneity of fat intake within geographically confined populations. The aim of the present study was to investigate prospectively, in a large, geographically and culturally heterogeneous cohort of European women, the association between dietary fat and breast cancer risk.

Norway (EL, MK, and GS); the Department of Epidemiology, Catalan Institute of Oncology, Barcelona, Spain (CAG); the Public Health Institute of Navarra, and El Centro de Investigación Biomédica en Red (CIBER) en Epidemiología y Salud Pública (CIBERESP), Pamplona, Spain (EA); the Public Health Division of Gipuzkoa, Donostia-San Sebastian, Spain (PA); the Epidemiology Department, Murcia Health Council, Murcia, and CIBERESP, Murcia, Spain (MJT); the Andalusian School of Public Health and CIBERESP, Granada, Spain (CM-G); the Health Information Unit, Public Health and Health Planning Directorate, Asturias, Spain (JRQ); the Department of Clinical Sciences, Lund University, Malmö, Sweden (GB and BG); the Department of Public Health and Clinical Medicine, Nutritional Research, Umeå University, Umeå, Sweden (GH); the Oncology Unit, Department of Radiation Sciences, Umeå University, Umeå, Sweden (PLe); the Center for Nutrition and Health, National Institute of Public Health and the Environment, Bilthoven, Netherlands (HBBdM and FJPvD); the Julius Center for Health Sciences and Primary Care, University Medical Center Utrecht, Utrecht, Netherlands (PHMP and CHvG); the Cancer Research UK Epidemiology Unit, University of Oxford, Oxford, United Kingdom (TKK and FLC); the Dunn Human Nutrition Unit, Medical Research Council, Cambridge, United Kingdom (SB and KTK); and the Department of Epidemiology & Public Health, Imperial College, London, United Kingdom (TN and ER).

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SUBJECTS AND METHODS

Subjects

The European Prospective Investigation into Cancer and Nutrition (EPIC) is a prospective study being conducted in 23 centers in 10 European countries: Norway, Sweden, Denmark, the United Kingdom, Germany, the Netherlands, France, Spain, Italy, and Greece. The design and methods of EPIC were described in detail elsewhere (22, 23). Briefly, the study recruited 519 978 volunteers (men and women), who completed dietary and lifestyle questionnaires and whose anthropometric measurements were recorded. Most participants were recruited from the general population of a town or province, with the following exceptions: the French cohort was recruited from female members of the teachers' health insurance organization; the Utrecht (Netherlands) and Florence (Italy) cohorts were recruited from women presenting for breast cancer screening; the Ragusa and Turin (Italy) cohort and part of the Spanish cohort were recruited from blood donors and their spouses. The Cambridge cohort and one-half of the Oxford cohort (United Kingdom) were recruited by general practitioners, and they formed the general population group. The rest of the Oxford cohort was recruited by mail in collaboration with vegetarian societies and was considered separately as a "health-conscious" group.

The study was conducted in 328 238 women aged 20–70 y after exclusion of those with cancer at recruitment and those who did not complete the dietary or lifestyle questionnaires ($n = 3320$). To reduce the effect of implausible extreme values on the analysis, we also excluded women in whom the ratio of total energy intake to basal metabolic rate (24) was at either extreme of the distribution (cutoffs were the first and last percentiles; $n = 6764$). We also excluded 8412 women with missing values for the potentially confounding variables of smoking and education. Thus, the analyses were conducted in 319 826 women, and the mean duration of follow-up was 8.8 y (2 812 609.6 person-years).

All participants provided written informed consent. The present study was approved by the ethics review boards of the International Agency for Research on Cancer (Lyon, France) and of all local recruiting institutions.

Ascertainment of cancer cases

Cases were ascertained by population-based cancer registries in 7 of the participating countries (Denmark, Italy, the Netherlands, Spain, Sweden, the United Kingdom, and Norway). In France, Germany, Greece, and the Italian center of Naples, various methods were used to identify cases, including consulting health insurance records and regional or national pathology registries, and active follow-up was conducted by contacting participants or next-of-kin. By February 2007, all centers had provided follow-up data for cancer cases up to December 1998, and more recent follow-up data (up to December 2005) was provided in many cases. The International Classification of Diseases for Oncology (ICD 10-O-2) was used to code cases. At the latest central audit (end of February 2007), 7298 invasive breast cancer cases had been identified (ICD code C50). After we applied the exclusions described above, a total of 7119 cases of invasive breast cancer were investigated.

Dietary measurements

Diet was assessed by using country-specific (or in some cases center-specific) dietary questionnaires designed to capture local dietary habits. Eight countries used self-administered dietary questionnaires, whereas, in Greece, Spain, and southern Italy (Naples and Ragusa), the questionnaires were administered by interviewers. In most countries, the questionnaires were extensive quantitative instruments (containing up to 260 food items). In Denmark, Norway, Sweden (Umeå), and Italy (Naples), semi-quantitative food-frequency questionnaires (FFQs) were administered. In Sweden (Malmö), an interview-based diet history method combining a questionnaire with a 7-d menu book was used. In the United Kingdom, an FFQ and a 7-d dietary record were used, but all results are from the FFQ (22, 23). All dietary questionnaires had been validated (25–32). For most centers, validation employed 12 monthly 24-h recall interviews. Pearson correlation coefficients between the mean consumption of total fat, reported in the dietary questionnaire, and the mean from the 24-h recalls varied, for women, from ≈ 0.4 to 0.7. For saturated, monounsaturated, and polyunsaturated fat, correlation coefficients were in the range of 0.5–0.7, 0.4–0.8, and 0.4–0.7, respectively (25–31). A separate standardized 24-h dietary recall (24-hDR) interview was carried out in a random sample (8%) of the entire EPIC cohort (33) to correct measurement errors in the dietary questionnaire; the random sample was stratified by sex and 5-y age strata in proportion to the expected numbers of total cancer cases in those categories. The EPIC Nutrient Database (34) was used to convert the quantities of food consumed into daily energy and total, saturated, monounsaturated, and polyunsaturated fat intakes.

Statistical analysis

We used multivariate Cox proportional hazard models to assess the association of fat and fat subtype intake with breast cancer risk, with stratification by center, to control for center effects and age (1-y categories). In all models, age was the primary time variable. Statistical adjustments were performed for menopausal status [premenopausal, perimenopausal, postmenopausal, or uncertain, as defined by Lahmann et al (35)], alcohol intake (none or <12 , 12–24, or >24 g/d), height, weight, smoking status (never, former, or current), and educational attainment (years of schooling). Information on family history of breast cancer was not available in all centers and was not included in the analyses. Age at first delivery, age at menarche, the number of full-term pregnancies, and oral contraceptive use (not available for 35 852 women including 882 cases) were included as potential confounders in preliminary models but had no influence on the results, and these data were excluded from final models.

Because macronutrient intake correlates strongly with energy intake, we used various models to adjust for the confounding effect of energy intake. We first applied a standard model that included absolute fat intake (in g) and total nonalcoholic energy intake (36). We next applied a residual model in which we regressed each type of fat intake on total nonalcohol energy by using country-specific regression models; the residual of the nutrient intake together with nonalcohol energy intake was then included in the model (36). Third, we applied a density model, in which energy from fat as a proportion of total nonalcohol energy intake, together with nonalcohol energy, was included. The

fourth model was an energy partition model (36), in which non-alcohol energy from the specific nutrient was modeled as one variable, and nonalcohol energy from other sources (eg, protein, carbohydrates, or other fat subtypes as appropriate) was the second energy variable. The first 3 models estimated the hazard ratio (HR) of breast cancer associated with isocaloric replacement of other macronutrients with the nutrient of interest (fat). The fourth model estimated the HR associated with the addition of energy from fat to the diet, while keeping the quantities of other macronutrients constant. We also used a fifth model, a multivariate nutrient density model, to evaluate the effect of replacing non-alcohol energy from carbohydrate (20% decrease) with energy from fat (20% increase) in an isocaloric diet with mutual adjustment for protein intake, nonalcohol energy, and other fat subtypes as appropriate.

Fat intakes were analyzed as both categorical and continuous variables. For the former, quintiles of fat intake were determined from the distribution in the whole cohort. Tests of linear trend were performed by modeling the median of each quintile. When intakes of total fat, fat subtype, and energy components were modeled as continuous variables, they were transformed to logarithms to the base 1.2, so that HRs represent the risk associated with a 20% increase in intake.

We also examined whether the association between fat and breast cancer risk was modified by body mass index (BMI; in kg/m^2) (≥ 25 or <25), baseline menopausal status (postmenopausal or premenopausal), and menopausal hormone replacement therapy (HRT) use (sometime HRT use or no use). This evaluation was achieved by modeling product terms of dichotomized variables multiplied by the median of the fat intake quintile to which the subject belonged or multiplied by the subject's fat intake when fat was considered a continuous variable. The significance of interactions was assessed by comparing the likelihood ratio test statistics of the models with and without the product term to a chi-square distribution with 1 df.

To correct for error in food intake measurement, we used a center-specific multivariate calibration approach (37), in which reference values of each fat type (according to the type of energy adjustment model used) and nonalcohol energy, obtained from the 24-hDR, were regressed on the dietary questionnaire measurements, and the resulting coefficients were used to estimate intake values. Adjustments for weight, height, age at recruitment, alcohol consumption categories, smoking status, educational attainment, and study center were included. In addition, the models were adjusted for unequal seasonal and weekday distribution (weekday or weekend) of the 24-hDR interviews, and the ratio of the expected to the observed number of interviews at each center was applied as weighting.

The Q test statistic with 9 df was used to assess statistical heterogeneity and investigate the hypothesis that associations between dietary components and breast cancer risk were the same in all countries (38). We tested the proportional hazard assumption for each fat and fat subtype in relation to breast cancer risk by using the method of Grambsch and Therneau (39). In all cases, the proportional hazards assumption was satisfied for total fat and fat subtypes ($P > 0.22$).

All analyses were performed with STATA software (version 7.0; Stata Corp, College Station, TX). All tests were 2-tailed, and 95% CIs were calculated and used to determine statistical

TABLE 1

Daily energy intake from fat and fat subtypes in women participating in the European Prospective Investigation into Cancer and Nutrition

| Country | Cases | Person-years of follow-up | Total fat | Saturated fat | Monounsaturated fat | Polyunsaturated fat |
|------------------------|----------|---------------------------|--------------------------|---------------|---------------------|---------------------|
| | <i>n</i> | <i>n</i> | % | % | % | % |
| Denmark | 822 | 214 842 | 33.3 ± 5.22 [†] | 13.1 ± 2.76 | 11.1 ± 2.04 | 5.5 ± 1.48 |
| France | 2272 | 691 567 | 36.7 ± 5.82 | 14.4 ± 3.13 | 12.9 ± 2.54 | 6.6 ± 2.23 |
| Greece | 103 | 93 832 | 46.9 ± 4.73 | 13.1 ± 2.57 | 22.7 ± 4.11 | 7.0 ± 2.98 |
| Germany | 457 | 226 510 | 35.5 ± 5.54 | 14.8 ± 2.91 | 12.2 ± 2.12 | 6.2 ± 1.75 |
| Italy | 670 | 255 538 | 35.3 ± 5.58 | 12.4 ± 2.54 | 16.6 ± 3.40) | 4.3 ± 1.22 |
| Netherlands | 567 | 227 662 | 35.9 ± 5.45 | 15.3 ± 2.81 | 12.0 ± 2.29 | 7.0 ± 1.80 |
| Norway | 441 | 198 456 | 34.9 ± 4.83 | 13.7 ± 2.46 | 11.1 ± 1.85 | 6.5 ± 1.56 |
| Spain | 319 | 240 836 | 37.6 ± 6.06 | 11.6 ± 3.08 | 15.7 ± 3.60 | 6.0 ± 2.33 |
| Sweden | 648 | 257 185 | 35.8 ± 6.07 | 15.5 ± 3.54 | 12.6 ± 2.24 | 5.4 ± 1.56 |
| United Kingdom | | | | | | |
| General population | 423 | 130 141 | 33.5 ± 6.07 | 12.6 ± 3.24 | 11.1 ± 2.17 | 7.1 ± 2.19 |
| Health-conscious group | 397 | 276 038 | 32.3 ± 6.54 | 11.6 ± 3.39 | 10.4 ± 2.34 | 7.6 ± 2.44 |
| Total | 7119 | 2 812 609 | 35.7 ± 6.31 | 13.7 ± 3.25 | 13.0 ± 3.76 | 6.3 ± 2.17 |

[†] $\bar{x} \pm \text{SD}$ (all such values).

significance; interactions and heterogeneity were considered significant at $P < 0.05$.

RESULTS

The mean daily proportion of nonalcohol energy intake from total fat and fat subtypes obtained from the dietary questionnaires, by country, is shown in **Table 1**. Mean energy from total fat ranged from 32.2% in the UK health-conscious group to 46.9% in Greece; mean energy from saturated fats ranged from 11.6% in Spain and the UK health-conscious group to 15.5% in Sweden and the Netherlands. Mean energy from monounsaturated fats was highest in Greece (22.7%) and lowest, less than half that (10.4%), in the UK health-conscious group. Mean energy from polyunsaturated fat ranged from 4.3% in Italy to >7.6% in the United Kingdom.

Adjusted HRs for breast cancer by intake of total fat, obtained by using the 4 energy-adjustment models, are shown in **Table 2**. Irrespective of whether total fat intake was analyzed as a categorical or continuous variable, HRs indicated no statistically significant association of total fat intake with breast cancer risk. No significant between-country heterogeneity was found in any of these models.

Adjusted HRs of developing breast cancer by quintile of saturated, monounsaturated, or polyunsaturated fat for each of the 4 energy-adjustment models are shown in **Table 3**. Women in the highest quintile of saturated fat intake had a significantly greater risk of developing breast cancer than did those in the lowest quintile, according to the standard (HR: 1.13; 95% CI: 1.00, 1.27), density (HR: 1.10; 95% CI: 1.01, 1.19), and partition (1.11; 95% CI: 1.00, 1.23) models. The test for trend was statistically significant for the standard ($P = 0.038$) and partition ($P = 0.045$) models. No significant between-country heterogeneity (Q test) was found in any of these models ($P = 0.904, 0.761, 0.902$, and 0.927 for standard, residual, density, and partition models, respectively).

When fat intake was considered as a continuous variable, an increase in saturated fat intake was associated with greater risk in all models (HR: 1.02; 95% CI: 1.00, 1.04 for the standard and

density models; HR: 1.02; 95% CI: 1.00, 1.03 for the residual and partition models). After calibration, HRs for saturated fat intake remained statistically significant in most models and were slightly greater than those from the continuous model with the use of noncalibrated data.

No association of the other types of fat intake with breast cancer risk was observed, although HR estimates for monounsaturated fat were in the same direction as those for saturated fats, whereas HR estimates for polyunsaturated fat tended to go in the opposite direction (ie, were <1). After calibration, HRs for monounsaturated fat intake became statistically significant in most models.

When all types of fat were included simultaneously in the model together with protein intake (multivariate density model) to estimate the effect of isocalorically substituting the intake of a given fat subtype for carbohydrate, HRs indicated no statistically significant association between fat intake of any kind and breast cancer risk; however, the trends for saturated and monounsaturated fat remained (data not shown). We conducted a further subanalysis to exclude early cases (<2-y follow-up) because of the possible influence of subclinical disease on dietary fat intake, but the results did not change.

The association between fat consumption and breast cancer was not modified by BMI or menopausal status (data not shown). However, among menopausal women, the results obtained with the standard model differed according to reported HRT use (**Table 4**). Specifically, tests for an HRT × fat interaction were statistically significant for intakes of total fat ($P = 0.039$), monounsaturated fat ($P = 0.011$), and saturated fat ($P = 0.018$) but not for polyunsaturated fat ($P = 0.763$). Among women who never used HRT, greater breast cancer risk was associated with greater consumption of saturated fat (HR: 1.03; 95% CI: 1.00, 1.07 for 20% higher use). Among HT users, there was no association between saturated fat intake and breast cancer risk (HR: 0.99; 95% CI: 0.96, 1.03). Similar results were obtained for total and monounsaturated fat, but none were statistically significant. The other energy-adjustment methods gave similar results for all types of fat (data not shown). No significant between-country heterogeneity was found for the interaction between HRT and fat subtypes.

TABLE 2

Multivariate-adjusted hazard ratios (HRs) (and 95% CIs) for invasive breast cancer risk in association with total fat intake by quintile (Q) in the European Prospective Investigation into Cancer and Nutrition¹

| Energy adjustment model | Intake | | | | | <i>P</i> for trend ² | Intake as a continuous variable ³ | Calibrated data (Intake as continuous variable) ⁴ |
|---------------------------------|--------------|--------------------------------|-------------------|-------------------|-------------------|---------------------------------|----------------------------------------------|--------------------------------------------------------------|
| | Q1 | Q2 | Q3 | Q4 | Q5 | | | |
| Standard ⁵ | | | | | | | | |
| Median (g/d) | 46.0 | 61.3 | 74.0 | 88.6 | 113.4 | | | |
| Cases/person-years (<i>n</i>) | 1274/534 091 | 1375/547 764 | 1444/561 608 | 1522/576 957 | 1504/592 190 | | | |
| HR | 1 | 1.01 (0.93, 1.10) ⁶ | 1.03 (0.93, 1.13) | 1.05 (0.94, 1.17) | 1.02 (0.90, 1.17) | 0.601 | 1.02 (0.99, 1.04) | 1.04 (0.98, 1.10) |
| Residual ⁷ | | | | | | | | |
| Median (g/d) | 61.5 | 71.2 | 77.2 | 83.4 | 94.0 | | | |
| Cases/person-years (<i>n</i>) | 1352/559 634 | 1473/551 288 | 1432/553 818 | 1429/568 379 | 1433/579 490 | | | |
| HR | 1 | 1.10 (1.01, 1.19) | 1.07 (0.99, 1.17) | 1.06 (0.98, 1.16) | 1.06 (0.97, 1.16) | 0.378 | 1.02 (0.99, 1.04) | 1.03 (0.98, 1.09) |
| Density ⁸ | | | | | | | | |
| Median (% of energy/d) | 28.9 | 33.5 | 36.7 | 39.8 | 44.9 | | | |
| Cases/person-years (<i>n</i>) | 1347/548 716 | 1391/553 292 | 1516/559 319 | 1439/573 902 | 1426/577 379 | | | |
| HR | 1 | 1.00 (0.93, 1.08) | 1.07 (0.99, 1.15) | 0.99 (0.92, 1.07) | 1.04 (0.96, 1.13) | 0.432 | 1.02 (0.99, 1.04) | 1.04 (0.98, 1.09) |
| Partition ⁹ | | | | | | | | |
| Median (g/d) | 46.0 | 61.3 | 74.0 | 88.6 | 113.4 | | | |
| Cases/person-years (<i>n</i>) | 1274/534 090 | 1375/547 764 | 1444/561 616 | 1522/576 960 | 1504/592 178 | | | |
| HR | 1 | 1.01 (0.93, 1.09) | 1.02 (0.94, 1.11) | 1.04 (0.96, 1.14) | 1.02 (0.92, 1.12) | 0.575 | 1.01 (0.99, 1.03) | 1.03 (0.99, 1.07) |

¹ Values were stratified by age and center and adjusted for educational attainment, smoking status, height, weight, alcohol intake, and menopausal status. *n* = 7119 breast cancer cases.

² Test for linear trend was performed on median intake for each quintile.

³ Values were log_{1.2} transformed (equivalent to an increase of 20%).

⁴ Calibrated data were obtained by linear regression models that compared observed nutrient questionnaire measurements with 24-h dietary recall.

⁵ Standard models contain log-transformed total fat intake and log-transformed nonalcohol energy intake.

⁶ 95% CIs in parentheses (all such values).

⁷ Residual models contain the residual of the regression of log-transformed total fat intake on log-transformed nonalcohol energy intake and log-transformed nonalcohol energy intake.

⁸ Density models contain log-transformed nonalcohol energy intake from total fat and log-transformed nonalcohol energy intake.

⁹ Partition models contain log-transformed total fat and log-transformed nonalcohol energy intake.

DISCUSSION

In this prospective study in 319 826 women, 7119 of whom developed invasive breast cancer, we found a weak but significant association of breast cancer risk with dietary saturated fat intake but no association with total fat intake. Monounsaturated fat was positively associated with breast cancer incidence; HRs were similar to those for saturated fat and were statistically significant only after calibration. Previous observational studies, reviewed in the Introduction, provided conflicting evidence of such an association. The randomized Women's Health Initiative Dietary Modification Trial (16) suggested that reducing fat intake may lower breast cancer risk: the incidence was 9% lower (not significant) in the low-fat dietary pattern group than in control subjects after 8.1 y of follow-up. Furthermore, the most recent meta-analysis of case-control and cohort studies found a statistically significant relation between high saturated fat intake and greater breast cancer risk (15). The present analysis considered 12 cohort studies and 3783 cancer cases. The EPIC study included nearly twice as many breast cancer cases and achieved similar results. The recently published National Institutes of Health–AARP study of a large cohort of postmenopausal US

women found a direct association between breast cancer risk and saturated fat consumption (13). That study also found that intakes of total, monounsaturated, and polyunsaturated fat were significantly related to breast cancer risk, although only saturated fat remained significantly associated when all fat subtypes were mutually adjusted. The lack of a significant association between total fat and breast cancer risk in the present study could be due to the fact that saturated (and to some extent monounsaturated) fat directly increases breast cancer risk, whereas polyunsaturated fat is inversely related to that risk. When fat subtypes were mutually adjusted, saturated and monounsaturated fat consumption remained directly but not significantly associated with risk. In this mutual adjustment (ie, multivariate density model), analysis of fat type is complicated because the main dietary sources of monounsaturated fat and saturated fat are often the same, so it is difficult to separate the effect of one from another. Nevertheless, the results of this model support direct effects of both saturated and monounsaturated fats on risk. A direct effect of monounsaturated intake on breast cancer risk is supported by some studies (13, 20, 40), although others found no association (9, 11, 12) or an inverse association (17–19).

TABLE 3
Multivariate-adjusted hazard ratios (HRs) (and 95% CIs) of invasive breast cancer risk associated with subtypes of fat intake by quintile (Q) in the cohort of the European Prospective Investigation into Cancer and Nutrition¹

| Fat subtype | Intake | | | | | P for trend ² | Intake as continuous variable ³ | Calibrated data (intake as continuous variable) ⁴ |
|------------------------|--------------|--------------------------------|-------------------|-------------------|-------------------|--------------------------|--------------------------------------------|--------------------------------------------------------------|
| | Q1 | Q2 | Q3 | Q4 | Q5 | | | |
| Saturated fat | | | | | | | | |
| Standard ⁵ | | | | | | | | |
| Median (g/d) | 16.2 | 22.6 | 27.9 | 34.1 | 45.0 | | | |
| Cases/person-years (n) | 1160/538 852 | 1341/545 846 | 1451/557 165 | 1536/571 949 | 1631/598 798 | | | |
| HR | 1 | 1.08 (0.99, 1.17) ⁶ | 1.12 (1.02, 1.22) | 1.14 (1.03, 1.26) | 1.13 (1.00, 1.27) | 0.038 | 1.02 (1.00, 1.04) | 1.03 (1.00, 1.07) |
| Residual ⁷ | | | | | | | | |
| Median (g/d) | 21.1 | 26.1 | 29.3 | 32.6 | 38.1 | | | |
| Cases/person-years (n) | 1274/565 391 | 1342/547 796 | 1370/547 072 | 1454/560 043 | 1679/592 308 | | | |
| HR | 1 | 1.06 (0.98, 1.15) | 1.06 (0.97, 1.15) | 1.05 (0.97, 1.15) | 1.06 (0.98, 1.16) | 0.233 | 1.02 (1.00, 1.03) | 1.03 (0.99, 1.07) |
| Density ⁸ | | | | | | | | |
| Median (% of energy/d) | 9.9 | 12.3 | 13.9 | 15.5 | 18.2 | | | |
| Cases/person-years (n) | 1185/555 735 | 1374/547 967 | 1412/552 679 | 1474/565 272 | 1674/590 956 | | | |
| HR | 1 | 1.07 (0.99, 1.16) | 1.06 (0.98, 1.15) | 1.05 (0.96, 1.13) | 1.10 (1.01, 1.19) | 0.068 | 1.02 (1.00, 1.04) | 1.04 (1.00, 1.07) |
| Partition ⁹ | | | | | | | | |
| Median (kcal/d) | 145.8 | 203.3 | 250.9 | 306.5 | 404.6 | | | |
| Cases/person-years (n) | 1160/538 852 | 1341/545 852 | 1451/557 159 | 1536/571 961 | 1631/598 786 | | | |
| HR | 1 | 1.07 (0.98, 1.16) | 1.11 (1.02, 1.21) | 1.12 (1.02, 1.23) | 1.11 (1.00, 1.23) | 0.045 | 1.02 (1.00, 1.03) | 1.03 (1.00, 1.06) |
| Monounsaturated fat | | | | | | | | |
| Standard ⁵ | | | | | | | | |
| Median (g/d) | 15.3 | 21.0 | 26.0 | 32.3 | 44.1 | | | |
| Cases/person-years (n) | 1273/525 283 | 1352/546 266 | 1515/568 799 | 1584/58 914 | 1395/583 116 | | | |
| HR | 1 | 0.98 (0.91, 1.07) | 1.05 (0.95, 1.15) | 1.07 (0.96, 1.18) | 1.05 (0.92, 1.20) | 0.254 | 1.02 (0.99, 1.04) | 1.05 (1.00, 1.10) |
| Residual ⁷ | | | | | | | | |
| Median (g/d) | 20.2 | 24.2 | 27.0 | 30.3 | 38.1 | | | |
| Cases/person-years (n) | 1310/536 795 | 1469/551 560 | 1499/569 583 | 1552/586 099 | 1289/568 573 | | | |
| HR | 1 | 1.09 (1.00, 1.18) | 1.03 (0.95, 1.12) | 1.10 (1.00, 1.20) | 1.08 (0.98, 1.21) | 0.150 | 1.02 (0.99, 1.04) | 1.04 (0.99, 1.10) |
| Density ⁸ | | | | | | | | |
| Median (% of energy/d) | 9.5 | 11.3 | 12.7 | 14.5 | 18.2 | | | |
| Cases/person-years (n) | 1291/532 384 | 1472/548 551 | 1483/570 103 | 1579/590 546 | 1294/571 026 | | | |
| HR | 1 | 1.07 (0.99, 1.15) | 1.02 (0.95, 1.11) | 1.06 (0.98, 1.15) | 1.05 (0.96, 1.16) | 0.323 | 1.02 (0.99, 1.04) | 1.05 (1.00, 1.10) |
| Partition ⁹ | | | | | | | | |
| Median (kcal/d) | 137.9 | 189.0 | 234.4 | 291.0 | 397.2 | | | |
| Cases/person-years (n) | 1273/525 283 | 1352/546 266 | 1515/568 799 | 1584/589 146 | 1395/583 116 | | | |
| HR | 1 | 0.98 (0.90, 1.06) | 1.04 (0.95, 1.14) | 1.06 (0.96, 1.17) | 1.04 (0.93, 1.18) | 0.253 | 1.01 (0.99, 1.03) | 1.04 (1.00, 1.09) |
| Polyunsaturated fat | | | | | | | | |
| Standard ⁵ | | | | | | | | |
| Median (g/d) | 7.2 | 9.8 | 12.3 | 15.4 | 21.3 | | | |
| Cases/person-years (n) | 1403/550 518 | 1409/547 583 | 1376/555 864 | 1445/570 597 | 1486/588 046 | | | |
| HR | 1 | 1.00 (0.93, 1.08) | 0.95 (0.88, 1.04) | 0.95 (0.87, 1.04) | 0.97 (0.88, 1.07) | 0.372 | 0.99 (0.98, 1.01) | 0.99 (0.95, 1.03) |
| Residual ⁷ | | | | | | | | |
| Median (g/d) | 8.4 | 10.9 | 12.7 | 15.0 | 19.5 | | | |
| Cases/person-years (n) | 1523/572 752 | 1390/556 567 | 1377/552 245 | 1470/559 334 | 1359/571 712 | | | |
| HR | 1 | 0.97 (0.90, 1.06) | 1.02 (0.94, 1.10) | 1.00 (0.92, 1.08) | 0.99 (0.91, 1.08) | 0.967 | 0.99 (0.98, 1.01) | 0.99 (0.96, 1.03) |
| Density ⁸ | | | | | | | | |
| Median (% of energy/d) | 4.0 | 5.1 | 6.1 | 7.2 | 9.4 | | | |
| Cases/person-years (n) | 1533/573 440 | 1389/557 490 | 1374/552 199 | 1472/558 845 | 1351/570 636 | | | |
| HR | 1 | 0.99 (0.91, 1.07) | 0.95 (0.88, 1.03) | 0.99 (0.92, 1.07) | 0.96 (0.88, 1.04) | 0.390 | 0.99 (0.98, 1.01) | 0.99 (0.95, 1.02) |
| Partition ⁹ | | | | | | | | |
| Median (kcal/d) | 64.4 | 88.6 | 110.7 | 138.6 | 191.5 | | | |
| Cases/person-years (n) | 1403/550 518 | 1409/547 583 | 1376/555 864 | 1445/570 598 | 1486/588 046 | | | |
| HR | 1 | 1.00 (0.93, 1.08) | 0.95 (0.88, 1.04) | 0.95 (0.87, 1.04) | 0.97 (0.88, 1.07) | 0.370 | 0.99 (0.98, 1.01) | 0.99 (0.95, 1.02) |

¹ Values were stratified by age and center and were adjusted for educational attainment, smoking status, height, weight, alcohol intake, and menopausal status. *n* = 7119 breast cancer cases.

² Test for linear trend was performed on median intake for each quintile.

³ Values were log_{1,2} transformed (equivalent to a 20% increase).

⁴ Calibrated data were obtained by linear regression models that compare observed nutrient questionnaire measurements with 24-h dietary recall.

⁵ Standard models contain log-transformed fat subtype intake and nonalcohol energy intake.

⁶ 95% CIs in parentheses (all such values).

⁷ Residual models contain residual of regression of log-transformed fat subtype intake and nonalcohol energy intake.

⁸ Density models contain log-transformed energy from fat subtypes and nonalcohol energy.

⁹ Partition models contain log-transformed fat subtype intake and energy intake from other sources excluding alcohol.

TABLE 4

Multivariate-adjusted hazard ratios (HRs) (and 95% CIs) for invasive breast cancer risk in association with subtypes of fat intake by quintile (Q) in postmenopausal women stratified by hormone replacement therapy (HRT)¹

| Subtypes of fat and quintile of intake | HRT (<i>n</i> = 1909) | No HRT (<i>n</i> = 1553) | <i>P</i> for interaction |
|----------------------------------------|--------------------------------|---------------------------|--------------------------|
| Total fat | | | |
| Q1 | 1 | 1 | |
| Q2 | 0.91 (0.78, 1.06) ² | 1.12 (0.95, 1.32) | |
| Q3 | 1.01 (0.86, 1.19) | 0.98 (0.82, 1.17) | |
| Q4 | 0.87 (0.73, 1.05) | 1.02 (0.84, 1.23) | |
| Q5 | 0.85 (0.69, 1.05) | 1.09 (0.88, 1.36) | |
| <i>P</i> for trend ³ | 0.127 | 0.74 | 0.053 |
| Intake as continuous variable | 0.98 (0.94, 1.02) | 1.02 (0.98, 1.06) | 0.039 |
| Saturated fat | | | |
| Q1 | 1 | 1 | |
| Q2 | 1.14 (0.98, 1.34) | 0.98 (0.83, 1.16) | |
| Q3 | 1.11 (0.95, 1.31) | 1.10 (0.92, 1.30) | |
| Q4 | 1.02 (0.86, 1.22) | 1.08 (0.90, 1.30) | |
| Q5 | 1.01 (0.83, 1.23) | 1.21 (0.99, 1.48) | |
| <i>P</i> for trend ³ | 0.698 | 0.044 | 0.020 |
| Intake as a continuous variable | 0.99 (0.96, 1.03) | 1.03 (1.00, 1.07) | 0.018 |
| Monounsaturated fat | | | |
| Q1 | 1 | 1 | |
| Q2 | 0.92 (0.79, 1.07) | 1.09 (0.92, 1.29) | |
| Q3 | 0.98 (0.84, 1.15) | 1.09 (0.92, 1.31) | |
| Q4 | 0.94 (0.79, 1.12) | 1.06 (0.88, 1.29) | |
| Q5 | 0.90 (0.73, 1.11) | 1.17 (0.94, 1.46) | |
| <i>P</i> for trend ³ | 0.426 | 0.239 | 0.05 |
| Intake as a continuous variable | 0.98 (0.95, 1.02) | 1.03 (0.99, 1.06) | 0.011 |
| Polyunsaturated fat | | | |
| Q1 | 1 | 1 | |
| Q2 | 1.01 (0.87, 1.19) | 1.13 (0.96, 1.32) | |
| Q3 | 0.91 (0.78, 1.07) | 1.02 (0.86, 1.21) | |
| Q4 | 0.90 (0.76, 1.06) | 0.99 (0.83, 1.18) | |
| Q5 | 0.97 (0.82, 1.16) | 0.95 (0.79, 1.15) | |
| <i>P</i> for trend ³ | 0.460 | 0.298 | 0.763 |
| Intake as a continuous variable | 0.99 (0.97, 1.02) | 0.99 (0.96, 1.01) | 0.808 |

¹ The standard model was used. Values were stratified by age and center and were adjusted for educational attainment, smoking status, height, weight, alcohol intake, and menopausal status. Likelihood ratio test on the median intake level in each quintile with 1 df and likelihood ratio test on the continuous scale with 1 df.

² HR; 95% CI in parentheses (all such values).

³ Test for linear trend was performed by using the median intake in each quintile.

Various biological data support a relation between high fat intake and greater breast cancer risk. Why, then, have many studies found no association of fat with breast cancer? One possible explanation is that the populations investigated were characterized by uniform dietary fat intake. In the EPIC study, which was designed to obtain data from populations that varied considerably in quantity of fat consumed, the interquintile range for the entire cohort was 54.5–98.3 g/d. Intakes of saturated and monounsaturated fats (expressed as % of energy) also varied markedly across cohorts: in Greece, Italy, and Spain, more calories were obtained from monounsaturated than saturated fat, whereas, in all other countries, more calories were obtained from saturated than monounsaturated fat.

A second possible explanation is that there is no generally agreed-upon method of adjustment for the influence of total energy intake (which is closely related to total macronutrient intake). Various energy-adjustment methods have been used, including the multivariate nutrient density (8, 9, 13, 41), residual (13, 42–45), energy partition (13, 45–47), and nutrient density (8, 13, 41) methods. We used 4 adjustment methods (ie, standard, residual, density, and energy partition), which are similar to those

used in the National Institutes of Health–AARP study (13). In that study, fat intake was found to be associated with breast cancer risk irrespective of the type of energy adjustment used. In the present study, when exposures were considered as continuous variables, all models gave similar results. These considerations suggest that the type of energy adjustment does not influence the detection or nondetection of an effect.

A third possible explanation is that errors in measuring dietary fat and energy intake obscure a fat-cancer association. This possibility is supported by that fact that 2 studies that assessed fat intake by using 2 independent methods found a stronger association of fat with breast cancer with a measurement method that was potentially more precise than the FFQ (48, 49). Unfortunately, the calibration procedure used by EPIC does not completely solve the problem of measurement error, because errors inherent in the calibration instrument (24-hDR) and the dietary questionnaire are not completely independent of each other (50). Nonetheless, after calibration, HRs for saturated fat intake were slightly greater than uncalibrated data when intake was a continuous variable (the only situation in which calibration can be applied).

Despite inconsistencies across studies, ≥ 2 biological mechanisms lend support to the dietary fat–breast cancer hypothesis. First, fat intake may raise endogenous estrogen concentrations (51). Intervention studies indicate that reduced fat intake lowers the concentrations and bioavailability of serum sex hormones (52–54). Most established epidemiologic risk factors for breast cancer are related to alterations in hormone metabolism, and high concentrations of circulating sex hormones increase the risk of breast cancer (55, 56). In the present study, sex hormone involvement was suggested by significant interactions of the intakes of total, monounsaturated, and saturated fats with HRT use in postmenopausal women. Similar associations were reported in the National Institutes of Health–AARP cohort (13).

Other studies (57, 58) have also found that HRT influences the association between body fat (as BMI or adult weight gain) and breast cancer risk: the risk was increased only in obese postmenopausal women not using HRT. An earlier meta-analysis of 9 prospective studies also found a significant association between serum concentrations of estrogens and postmenopausal breast cancer, which, again, was confined to HRT nonusers at the time of blood donation (59). HRT use is such a strong risk factor for breast cancer (60) that it may conceal the effect of other factors that influence breast cancer risk by altering endogenous hormones. In particular, the increase in sex hormone concentrations due to high fat intake may be small compared with the already high concentrations due to HRT, and a high-fat diet is unlikely to confer additional risk for HRT users. In contrast, high fat intake in postmenopausal women not using HRT may increase estrogen bioavailability beyond the threshold that is likely to increase breast cancer risk. Consumption of fat, especially saturated fat, may also increase breast cancer risk by worsening insulin resistance (61). Several studies suggest associations of plasma concentrations of insulin, C peptide, and insulin growth factor-I with breast cancer (62). A previous EPIC report showed that high insulin growth factor-I and insulin growth factor–binding protein-3 were associated with a greater risk of breast cancer that was diagnosed after (but not before) age 50 y in women not using HRT or oral contraceptives at blood sampling (63).

Strengths of the present study are its large size (>7000 cases), prospective cohort design, wide variation in fat intake between centers, and extensive availability of information on potential confounders. Limitations common to most observational dietary studies are also apparent. In particular, the estimation of food and nutrient intake by questionnaires is associated with large random error that tends to attenuate relative risk estimates.

In conclusion, the data from the present study add to the accumulating evidence of a weak positive association between saturated fat intake and breast cancer risk. This association was more pronounced in postmenopausal women who never used HRT. It seems that a large cohort characterized by wide variation in fat consumption is necessary to show this weak effect, which may have limited significance from the public health point of view.

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DP, AO, ATj, M-CB-R, GS, ATr, AN, MUJ, PL, GH, JL, PA, SP, TJK, KTK, PL, RT, PV, EA, CHvG, MJT, CM-G, and BG: acquisition of data; SS, VK, VP, ACMT, FLC, KO, CAG, GB, MJ, VC, FC-C, TJK, ATj, UN, SR, and FJBvD: analysis and interpretation of data; and all authors: contributions to and review of the manuscript. None of the authors had a personal or financial conflict of interest. The sponsors had no role in study design, data collection, analysis and interpretation of results, or the writing of the manuscript.

REFERENCES

1. Key TJ, Verkasalo PK, Banks E. Epidemiology of breast cancer. *Lancet Oncol* 2001;2:133–40.
2. Armstrong B, Doll R. Environmental factors and cancer incidence and mortality in different countries, with special reference to dietary practices. *Int J Cancer* 1975;15:617–31.
3. Aylsworth CF, Jone C, Trosko JE, Meites J, Welsch CW. Promotion of 7,12-dimethylbenz[a]anthracene-induced mammary tumorigenesis by high dietary fat in the rat: possible role of intercellular communication. *J Natl Cancer Inst* 1984;72:637–45.
4. Prentice RL, Kakar F, Hursting S, Sheppard L, Klein R, Kushi LH. Aspects of the rationale for the Women's Health Trial. *J Natl Cancer Inst* 1988;80:802–14.
5. Rogers AE, Zeisel SH, Groopman J. Diet and carcinogenesis. *Carcinogenesis* 1993;14:2205–17.
6. Welsch CW. Relationship between dietary fat and experimental mammary tumorigenesis: a review and critique. *Cancer Res* 1992;52:2040s–8s.
7. Howe GR, Hirohata T, Hislop TG, et al. Dietary factors and risk of breast cancer: combined analysis of 12 case-control studies. *J Natl Cancer Inst* 1990;82:561–9.
8. Knekt P, Albanes D, Seppanen R, et al. Dietary fat and risk of breast cancer. *Am J Clin Nutr* 1990;52:903–8.
9. Velie E, Kulldorff M, Schairer C, Block G, Albanes D, Schatzkin A. Dietary fat, fat subtypes, and breast cancer in postmenopausal women: a prospective cohort study. *J Natl Cancer Inst* 2000;92:833–9.
10. Byrne C, Rockett H, Holmes MD. Dietary fat, fat subtypes, and breast cancer risk: lack of an association among postmenopausal women with no history of benign breast disease. *Cancer Epidemiol Biomarkers Prev* 2002;11:261–5.
11. Kim EH, Willett WC, Colditz GA, et al. Dietary fat and risk of postmenopausal breast cancer in a 20-year follow-up. *Am J Epidemiol* 2006;164:990–7.
12. Lof M, Sandin S, Lagiou P, et al. Dietary fat and breast cancer risk in the Swedish women's lifestyle and health cohort. *Br J Cancer* 2007;97:1570–6.
13. Thiebaut AC, Kipnis V, Chang SC, et al. Dietary fat and postmenopausal invasive breast cancer in the National Institutes of Health–AARP Diet and Health Study cohort. *J Natl Cancer Inst* 2007;99:451–62.
14. Smith-Warner SA, Spiegelman D, Adami HO, et al. Types of dietary fat and breast cancer: a pooled analysis of cohort studies. *Int J Cancer* 2001;92:767–74.
15. Boyd NF, Stone J, Vogt KN, Connelly BS, Martin LJ, Minkin S. Dietary fat and breast cancer risk revisited: a meta-analysis of the published literature. *Br J Cancer* 2003;89:1672–85.
16. Prentice RL, Caan B, Chlebowski RT, et al. Low-fat dietary pattern and risk of invasive breast cancer: the Women's Health Initiative Randomized Controlled Dietary Modification Trial. *JAMA* 2006;295:629–42.
17. Wolk A, Bergstrom R, Hunter D, et al. A prospective study of association of monounsaturated fat and other types of fat with risk of breast cancer. *Arch Intern Med* 1998;158:41–5.
18. Holmes MD, Hunter DJ, Colditz GA, et al. Association of dietary intake of fat and fatty acids with risk of breast cancer. *JAMA* 1999;281:914–20.
19. Voorrips LE, Brants HA, Kardiinal AF, Hiddink GJ, van den Brandt PA, Goldbohm RA. Intake of conjugated linoleic acid, fat, and other fatty acids in relation to postmenopausal breast cancer: the Netherlands Cohort Study on Diet and Cancer. *Am J Clin Nutr* 2002;76:873–82.
20. Cho ESDH. Premenopausal fat intake and risk of breast cancer. *J Natl Cancer Inst* 2003;95:1079–85.
21. Prentice RL. Measurement error and results from analytic epidemiology: dietary fat and breast cancer. *J Natl Cancer Inst* 1988;88:1738–47.
22. Riboli E, Kaaks R. The EPIC Project: rationale and study design. *European Prospective Investigation into Cancer and Nutrition. Int J Epidemiol* 1997;26(suppl):S6–14.

23. Riboli E, Hunt KJ, Slimani N, et al. European Prospective Investigation into Cancer and Nutrition (EPIC): study populations and data collection. *Public Health Nutr* 2002;5:1113–24.
24. Harris JA, Benedict FG. A biometric study of human basal metabolism. *Proc Natl Acad Sci U S A* 1918;4:370–3.
25. Ocke MC, Bueno-de-Mesquita HB, Pols MA, Smit HA, van Staveren WA, Kromhout D. The Dutch EPIC food frequency questionnaire. II. Relative validity and reproducibility for nutrients. *Int J Epidemiol* 1997;26(suppl):S49–58.
26. Bohlscheid-Thomas S, Hoting I, Boeing H, Wahrendorf J. Reproducibility and relative validity of energy and macronutrient intake of a food frequency questionnaire developed for the German part of the EPIC project. *European Prospective Investigation into Cancer and Nutrition. Int J Epidemiol* 1997;26(suppl):S71–81.
27. Relative validity and reproducibility of a diet history questionnaire in Spain. II. Nutrients. EPIC Group of Spain. *European Prospective Investigation into Cancer and Nutrition. Int J Epidemiol* 1997;26(suppl):S100–9.
28. van Liere MJ, Lucas F, Clavel F, Slimani N, Villeminot S. Relative validity and reproducibility of a French dietary history questionnaire. *Int J Epidemiol* 1997;26(suppl):S128–36.
29. Bingham SA, Gill C, Welch A, et al. Validation of dietary assessment methods in the UK arm of EPIC using weighed records, and 24-hour urinary nitrogen and potassium and serum vitamin C and carotenoids as biomarkers. *Int J Epidemiol* 1997;26(suppl):S137–51.
30. Pisani P, Faggiano F, Krogh V, Palli D, Vineis P, Berrino F. Relative validity and reproducibility of a food frequency dietary questionnaire for use in the Italian EPIC centres. *Int J Epidemiol* 1997;26(suppl):S152–60.
31. Riboli E, Elmstahl S, Saracci R, Gullberg B, Lindgarde F. The Malmo Food Study: validity of two dietary assessment methods for measuring nutrient intake. *Int J Epidemiol* 1997;26(suppl):S161–73.
32. Hjartaker A, Andersen LF, Lund E. Comparison of diet measures from a food-frequency questionnaire with measures from repeated 24-hour dietary recalls. *The Norwegian Women and Cancer Study. Public Health Nutr* 2007;10:1094–103.
33. Slimani N, Ferrari P, Ocke M, et al. Standardization of the 24-hour diet recall calibration method used in the European Prospective Investigation into Cancer and Nutrition (EPIC): general concepts and preliminary results. *Eur J Clin Nutr* 2000;54:900–17.
34. Slimani N, Deharveng G, Unwin I, et al. The EPIC Nutrient Database Project (ENDB): a first attempt to standardize nutrient databases across the 10 European countries participating in the EPIC study. *Eur J Clin Nutr* 2007;61:1037–56.
35. Lahmann PH, Hoffmann K, Allen N, et al. Body size and breast cancer risk: findings from the European Prospective Investigation into Cancer and Nutrition (EPIC). *Int J Cancer* 2004;111:762–71.
36. Willett W, Stampfer MJ. Total energy intake: implications for epidemiologic analyses. *Am J Epidemiol* 1986;124:17–27.
37. Rosner B, Spiegelman D, Willett WC. Correction of logistic regression relative risk estimates and confidence intervals for random within-person measurement error. *Am J Epidemiol* 1992;136:1400–13.
38. Cochran WG. The combination of estimate from different experiments. *Biometrics* 1954;10:101–29.
39. Grambsch PM, Therneau TM. Proportional hazards tests and diagnostics based on weighted residuals. *Biometrika* 1994;81:515–26.
40. Wirfalt E, Mattisson I, Gullberg B, Johansson U, Olsson H, Berglund G. Postmenopausal breast cancer is associated with high intakes of omega6 fatty acids (Sweden). *Cancer Causes Control* 2002;13:883–93.
41. Kushi LH, Sellers TA, Potter JD, et al. Dietary fat and postmenopausal breast cancer. *J Natl Cancer Inst* 1992;84:1092–9.
42. Jones DY, Schatzkin A, Green SB, et al. Dietary fat and breast cancer in the National Health and Nutrition Examination Survey I Epidemiologic Follow-up Study. *J Natl Cancer Inst* 1987;79:465–71.
43. van den Brandt PA, van't Veer P, Goldbohm RA, et al. A prospective cohort study on dietary fat and the risk of postmenopausal breast cancer. *Cancer Res* 1993;53:75–82.
44. Willett WC, Hunter DJ, Stampfer MJ, et al. Dietary fat and fiber in relation to risk of breast cancer. An 8-year follow-up. *JAMA* 1992;268:2037–44.
45. Kushi LH, Potter JD, Bostick RM, et al. Dietary fat and risk of breast cancer according to hormone receptor status. *Cancer Epidemiol Biomarkers Prev* 1995;4:11–9.
46. Graham S, Zielezny M, Marshall J, et al. Diet in the epidemiology of postmenopausal breast cancer in the New York State Cohort. *Am J Epidemiol* 1992;136:1327–37.
47. Howe GR, Friedenreich CM, Jain M, Miller AB. A cohort study of fat intake and risk of breast cancer. *J Natl Cancer Inst* 1991;83:336–40.
48. Bingham SA, Luben R, Welch A, Wareham N, Khaw KT, Day N. Are imprecise methods obscuring a relation between fat and breast cancer? *Lancet* 2003;362:212–4.
49. Freedman LS, Potischman N, Kipnis V, et al. A comparison of two dietary instruments for evaluating the fat-breast cancer relationship. *Int J Epidemiol* 2006;35:1011–21.
50. Kipnis V, Midthune D, Freedman LS, et al. Empirical evidence of correlated biases in dietary assessment instruments and its implications. *Am J Epidemiol* 2001;153:394–403.
51. Michels KB, Mohllajee AP, Roset-Bahmanyar E, Beehler GP, Moysich KB. Diet and breast cancer: a review of the prospective observational studies. *Cancer* 2007;109:2712–49.
52. Prentice RL, Sheppard L. Dietary fat and cancer: consistency of the epidemiologic data, and disease prevention that may follow from a practical reduction in fat consumption. *Cancer Causes Control* 1990;1:81–97.
53. Berrino F, Bellati C, Secreto G, et al. Reducing bioavailable sex hormones through a comprehensive change in diet: the Diet and Androgens (DIANA) randomized trial. *Cancer Epidemiol Biomarkers Prev* 2001;10:25–33.
54. Wu AH, Pike MC, Stram DO. Meta-analysis: dietary fat intake, serum estrogen levels, and the risk of breast cancer. *J Natl Cancer Inst* 1999;91:529–34.
55. Kaaks R, Berrino F, Key T, et al. Serum sex steroids in premenopausal women and breast cancer risk within the European Prospective Investigation into Cancer and Nutrition (EPIC). *J Natl Cancer Inst* 2005;97:755–65.
56. Kaaks R, Rinaldi S, Key TJ, et al. Postmenopausal serum androgens, oestrogens and breast cancer risk: the European Prospective Investigation into Cancer and Nutrition. *Endocr Relat Cancer* 2005;12:1071–82.
57. Beral V. Breast cancer and hormone-replacement therapy in the Million Women Study. *Lancet* 2003;362:419–27.
58. Breast cancer and hormone replacement therapy: collaborative reanalysis of data from 51 epidemiological studies of 52,705 women with breast cancer and 108,411 women without breast cancer. Collaborative Group on Hormonal Factors in Breast Cancer. *Lancet* 1997;350:1047–59.
59. Key T, Appleby P, Barnes I, Reeves G. Endogenous sex hormones and breast cancer in postmenopausal women: reanalysis of nine prospective studies. *J Natl Cancer Inst* 2002;94:606–16.
60. Campagnoli C, Clavel-Chapelon F, Kaaks R, Peris C, Berrino F. Progestins and progesterone in hormone replacement therapy and the risk of breast cancer. *J Steroid Biochem Mol Biol*. 2005;96:95–108.
61. Riccardi G, Giacco R, Rivellese AA. Dietary fat, insulin sensitivity and the metabolic syndrome. *Clin Nutr* 2004;23:447–56.
62. Kaaks R. Nutrition, insulin, IGF-1 metabolism and cancer risk: a summary of epidemiological evidence. *Novartis Found Symp* 2004;262:247–60.
63. Rinaldi S, Peeters PH, Berrino F, et al. IGF-I, IGFBP-3 and breast cancer risk in women: the European Prospective Investigation into Cancer and Nutrition (EPIC). *Endocr Relat Cancer* 2006;13:593–605.